

# On Wigner-Ville Spectra and the Unicity of Time-Varying Quantile-Based Spectral Densities

Stefan BIRR\*

Ruhr-Universität Bochum

Holger DETTE†

Ruhr-Universität Bochum

Marc HALLIN‡

ECARES, Université Libre de Bruxelles

Tobias KLEY §

London School of Economics

and

Stanislav VOLGUSHEV

University of Toronto

## Abstract

**Abstract:** The unicity of the time-varying quantile-based spectrum proposed in Birr et al. (2016) is established via an asymptotic representation result involving Wigner-Ville spectra.

**Keywords:** Copula-based spectrum, Laplace spectrum, Quantile-based spectrum, Time-varying spectrum, Wigner-Ville spectrum.

## 1 Introduction

Emmanuel Parzen is perhaps best known for his pioneering contributions to spectral methods in time series analysis and kernel density estimation—a method

---

\*Research supported by the Sonderforschungsbereich “Statistical modelling of nonlinear dynamic processes” (SFB 823, Teilprojekt A1, C1) of the Deutsche Forschungsgemeinschaft.

†Research supported by the Sonderforschungsbereich “Statistical modelling of nonlinear dynamic processes” (SFB 823, Teilprojekt A1, C1) of the Deutsche Forschungsgemeinschaft.

‡Research supported by an Interuniversity Attraction Pole (2012-2017) of the Belgian Science Policy Office.

§Supported by the Engineering and Physical Sciences Research Council grant no. EP/L014246/1.

which in the signal processing community is often referred to as the *Parzen-Rosenblatt window* method. Another lifelong interest of Emmanuel Parzen was quantile-based inference (see, for instance, Parzen (2004)). In 2009, when officially retiring from Texas A&M, he delivered a “Last Lecture”. A Last Lecture is supposed to convey the essential message of a scholarly career, and, quite significantly, Emmanuel Parzen chose to epitomize his sixty-year long activity as a researcher in statistics with the eloquent title “*Quantiles are Optimal*”.

When asked to participate in this memorial volume, we thought that there was no better way to pay tribute to Emmanuel Parzen’s memory and outstanding achievement than contributing something on quantile-based spectral analysis, at the intersection of his favorite subjects, quantiles and spectral analysis.

Quantile-based (equivalently, copula-based) spectral methods recently have attracted renewed attention in the time-series community. Pioneering work in this direction has been done by Hong (1999) and Li (2008), who coined the name *Laplace spectrum*. Similar ideas were further developed by Hagemann (2013), and extended into *cross-spectrum* and spectral *kernel* concepts by Dette et al. (2015) and Kley et al. (2016), where the asymptotic normality of smoothed periodogram-based estimators is established. Along with the stationary marginal distribution, those quantile-based spectra entirely characterize the bivariate distributions of all couples  $(X_t, X_{t-h})$ ,  $h = 1, 2, \dots$  in a stationary process  $\{X_t, t \in \mathbb{Z}\}$ , hence provide much more information than classical second-order spectra; computing them has been made possible by Kley (2016) and his R package ‘quantspec’. Additional contributions can be found in Li (2012, 2014), Lee and Rao (2012) and Davis et al. (2013).

Whether quantile-based or classical, spectral methods typically require long observation periods, and long observation periods, in general, imply that stationarity assumptions are violated. In Birr et al. (2016), we therefore introduce a locally stationary version of the concepts and methods initiated in Dette et al. (2015) and Kley et al. (2016). The type of local stationarity required in this context differs substantially from the usual notions proposed, for instance, by Dahlhaus (1997), Zhou and Wu (2009a and b), Roueff and von Sachs (2011) or Vogt (2012). Just as the latter, however, it involves the choice of families of approximating stationary processes—a choice which, at first sight, could have an impact on the resulting time-varying spectrum. In this paper, we show that this is not the case, and establish the unicity of the time-varying quantile-based spectra defined there, irrespective of the chosen family of approximating processes.

The proof of that unicity property relies on an asymptotic representation (Proposition 1) of the time-varying quantile-based spectra in terms of appropriate *Wigner-Ville spectra* that no longer involve any specific approximating stationary processes—an idea that, mutatis mutandis, goes back to Theorem 2.2 in Dahlhaus (1996).

The outline of the paper is as follows. In Section 2, we briefly introduce the concept of the Wigner-Ville spectrum. Section 3 similarly presents the locally stationary quantile-based spectra proposed in Birr et al. (2016). In Section 4, we establish the asymptotic representation result connecting those spectra with

the Wigner-Ville ones. The desired unicity follows as a corollary.

## 2 Wigner-Ville spectra.

The so-called *Wigner-Ville spectrum* for harmonizable processes (Martin and Flandrin 1985) originates in the signal processing literature, and constitutes one of the few links between the engineering and the statistical approaches to spectral analysis.

Spectral analysis in engineering and physics usually deals with (deterministic) signals  $s(t)$ . The spectrum of a signal, in that context, is seldom constant over time (think of music or speech), so that the concept of a time-varying spectrum or a *time-frequency representation* is quite essential. The problem is that no obvious definition for such a concept exists—the situation is somewhat similar to that in probability and mathematical statistics, where no obvious concept exists for the spectrum of a nonstationary process. As a result, different approaches have been used. The Wigner-Ville spectrum is based on one of the classical time-frequency representations of a complex continuous-time signal  $s(t)$ , called the *Wigner distribution*, of the form

$$\mathcal{W}(t, \omega) := \frac{1}{2\pi} \int s^*(t - \frac{1}{2}\tau) s(t + \frac{1}{2}\tau) e^{-i\tau\omega} d\tau, \quad (1)$$

where the star indicates complex conjugation. Eugene Wigner's original motivation (Wigner 1932) was to be able to calculate the quantum correction to the second virial coefficient of a gas, which indicates how it deviates from the *ideal gas* law. The Wigner distribution was later used, in connection with characteristic function methods, in the context of signal analysis by Ville (1948). One unpleasant feature of  $\mathcal{W}(t, \omega)$ , however, is that it can take negative values.

Now, given a nonstationary continuous-time process  $\{X(t)\}$ , we can define the covariance kernel as

$$R(s, t) := \text{Cov}(X(s), X(t)).$$

For a nonstationary process, this kernel is no longer a function of the lag  $|t-s|$ ; in order to define a local autocovariance function, one can use  $R(t - \tau/2, t + \tau/2)$ , which measures autocovariance at lag  $\tau$  between two process values centered about a timepoint  $t$ . A time-varying spectrum then can be obtained as the Fourier transform of  $R$  with respect to  $\tau$ , namely,

$$\Psi(t, \omega) := \frac{1}{2\pi} \int R(t - \tau/2, t + \tau/2) e^{-i\tau\omega} d\tau.$$

Spectra of this form have been considered, for example, by Mark (1970). But they share the same (main) disadvantage as the Wigner-Ville distribution: they are not necessarily positive. As a consequence, they could not compete with the theory of *evolutionary spectra* proposed by Priestley (1965) and the approach was not further pursued.

Results by Claassen and Mecklenbräuger (1980) and Flandrin and Escudid (1980) brought the attention of researchers back to the Wigner-Ville concepts. They independently proved that the Wigner-Ville distribution indeed enjoys almost all properties that are desirable for a time frequency representation. For a quick overview of these properties, see Hlawatsch and Boudreault-Bartels (1992) and, for a systematic approach, Loynes (1968). Moreover, some of the more important properties are not compatible with the requirement of a non-negative function, turning what was perceived as the main drawback of the Wigner-Ville spectrum into a necessary evil. Motivated by these results, Martin (1982) defines a time-varying spectrum, called the Wigner-Ville spectrum, based on the Wigner-Ville distribution (1).

A process  $\{X(t)\}$  is called *harmonizable* if its autocovariance function can be represented as

$$R(s, t) = \frac{1}{4\pi^2} \int \int f(\lambda, \mu) e^{i(\lambda s - \mu t)} d\lambda d\mu$$

for some function  $f$ . Harmonizable processes are a generalization of weakly stationary processes, and were used by Priestley as a starting point for his theory of evolutionary spectra. Martin (1982) defined the *Wigner-Ville spectrum* of a harmonizable process  $\{X(t)\}$  as

$$W(t, \omega) := \frac{1}{2\pi} \int f(\omega - \tau/2, \omega + \tau/2) e^{i\tau t} d\tau,$$

justifying his notation by showing that, under appropriate assumptions,

$$W(t, \omega) = E \left[ \int e^{-i\omega\tau} X(t + \tau/2) X^*(t - \tau/2) d\tau \right].$$

Martin and Flandrin (1985) extended this definition to discrete-time processes and proposed a class of estimators for it, based on a weighted covariance estimator

$$\hat{R}(t+k, t-k) := \sum_{s \in \mathbb{Z}} \Phi(m, 2k) X(s+m+k) X^*(s+m-k),$$

where  $\Phi$  is some adequate lag-window function. They derived the first and second moments for those estimators under the assumption of Gaussian processes, but, just as with the theory of evolutionary spectra, no meaningful asymptotic results could be derived.

The first theory allowing for asymptotic results in this nonstationary context was initiated some ten years later by Dahlhaus (1996) with the introduction of *locally stationary processes*. Dahlhaus considers triangular arrays of stationary processes  $\{X_{t,T}\}$  which are “close” to a family  $\{X_\theta(t)\}_{\theta \in (0,1)}$  whenever  $t/T$  is close to  $\theta$ , and defines the time-varying spectrum  $f(u, \omega)$  of  $\{X_{t,T}\}$  at rescaled time point  $u = t/T$  as the classical spectrum

$$f(u, \omega) := \frac{1}{2\pi} \sum_{k \in \mathbb{Z}} \text{Cov}(X_u(k), X_u(0)) e^{-i\omega k} \quad (2)$$

of the stationary process  $\{X_u(t)\}$ . Now, as the family of approximating processes  $X_u(t)$  is not unique, it is not obvious from the definition that the resulting spectrum is unique. But, considering the Wigner-Ville spectrum

$$f_T(u, \omega) := \sum_{s \in \mathbb{Z}} \text{Cov}(X_{[uT-s/2], T}, X_{[uT+s/2], T}) e^{-i\omega s},$$

Dahlhaus (1996) proves that, under appropriate assumptions,

$$\int_{-\pi}^{\pi} |f_T(u, \omega) - f(u, \omega)|^2 d\omega = o(1),$$

as  $T \rightarrow \infty$ , which, provided that  $\omega \mapsto f_T(u, \omega)$  and  $\omega \mapsto f(u, \omega)$  are continuous, implies unicity of  $\omega \mapsto f(u, \omega)$ .

### 3 Copula Spectral Densities for locally stationary processes

Consider a triangular array  $(X_{t,T}, 1 \leq t \leq T)$ ,  $T \in \mathbb{N}$ , of finite-length realizations of nonstationary processes  $\{X_{t,T}, t \in \mathbb{Z}\}$ ,  $T \in \mathbb{N}$ . The quantile-based or copula spectral density kernels of a stationary process are defined (Dette et al. (2015); Kley et al. (2016)) in terms of its bivariate marginal distribution functions. Therefore, it seems natural to use bivariate marginal distribution functions when evaluating, in the definition of local stationarity, the distance between the nonstationary process  $\{X_{t,T}\}$  and its stationary approximations  $\{X_u(t)\}$ . This leads to the following definition.

**Definition 1.** (Birr, Volgushev, Kley, Dette, and Hallin 2016). *A triangular array  $\{(X_{t,T})_{t \in \mathbb{Z}}\}_{T \in \mathbb{N}}$  of processes is called locally strictly stationary (of order two) if there exists a constant  $L > 0$  and, for every  $u \in (0, 1)$ , a strictly stationary process  $\{X_u(t), t \in \mathbb{Z}\}$  such that, for every  $1 \leq r, s \leq T$ ,*

$$\|F_{r,s;T}(\cdot, \cdot) - G_{r-s}^u(\cdot, \cdot)\|_\infty \leq L \left( \max(|r/T - u|, |s/T - u|) + 1/T \right), \quad (3)$$

where  $\|\cdot\|_\infty$  stands for the supremum norm, while  $F_{r,s;T}(\cdot, \cdot)$  and  $G_k^u(\cdot, \cdot)$  denote the joint distribution functions of  $(X_{r,T}, X_{s,T})$  and  $(X_u(0), X_u(-k))$ , respectively.

With this concept of nonstationarity, we can transfer the stationary quantile-based concepts of Dette et al. (2015) from  $\{X_u(t)\}$  to the nonstationary  $\{X_{t,T}\}$ , defining the *time-varying copula spectral density*. Define the *lag-h-copula cross-covariance kernel* of  $\{X_u(t)\}$  as

$$\gamma_h^u(\tau_1, \tau_2) := \text{Cov} \left( \mathbb{I}_{\{X_u(t) \leq q^u(\tau_1)\}}, \mathbb{I}_{\{X_u(t-h) \leq q^u(\tau_2)\}} \right), \quad \tau_1, \tau_2 \in (0, 1), \quad (4)$$

where  $q^u(\tau)$  denotes  $X_u(t)$ 's marginal quantile of order  $\tau$ . If we assume that for all  $\tau_1, \tau_2, u$  the lag-h-covariance kernels  $\gamma_h^u(\tau_1, \tau_2)$  are summable, we obtain

the *time-varying copula spectral density*

$$\mathfrak{f}^u(\omega, \tau_1, \tau_2) := \frac{1}{2\pi} \sum_{h=-\infty}^{\infty} \gamma_h^u(\tau_1, \tau_2) e^{-ih\omega}, \quad \tau_1, \tau_2 \in (0, 1), \quad \omega \in (-\pi, \pi]. \quad (5)$$

Comparing those definitions with those of the local spectral densities in Dahlhaus (1996), we see that the approximating autocovariances appearing there are replaced with autocovariances involving the copula transform (equivalently, the quantiles) of the approximating processes. This indicates that the local spectral density kernels (5) can be viewed as fully non-parametric generalizations of their classical  $L^2$ -based counterparts, capturing pairwise serial dependencies of arbitrary forms. For detailed comparisons, we refer to Dette et al. (2015) and Kley et al. (2016).

## 4 Unicity of the time-varying quantile-based spectral density

The Wigner-Ville spectrum associated with the series of indicators involved in the definition (3) of the copula cross-covariance kernel of  $X_u(t)$  takes the form

$$\mathfrak{W}_{t_0, T}(\omega, \tau_1, \tau_2) = \frac{1}{2\pi} \sum_{s=-\infty}^{\infty} \gamma_{t_0; T}(s, \tau_1, \tau_2) e^{-i\omega s}, \quad (6)$$

where

$$\gamma_{t_0; T}(s, \tau_1, \tau_2) := \text{Cov}\left(\mathbb{I}_{\{X_{\lfloor t_0 + s/2 \rfloor, T} \leq F_{\lfloor t_0 + s/2 \rfloor; T}^{-1}(\tau_1)\}}, \mathbb{I}_{\{X_{\lfloor t_0 - s/2 \rfloor, T} \leq F_{\lfloor t_0 - s/2 \rfloor; T}^{-1}(\tau_2)\}}\right)$$

and  $F_{t, T}^{-1}$  denotes the generalized inverse of  $F_{t, T}$ , the marginal distribution function of  $X_{t, T}$ . Indicators being bounded,  $\mathfrak{W}_{t_0, T}(\omega, \tau_1, \tau_2)$  exists and is uniquely defined for all  $\omega, \tau_1$  and  $\tau_2$  as soon as absolute summability holds for the right-hand side of (6). The following proposition establishes a strong relation between the Wigner-Ville spectrum (6) and the time-varying copula spectral density kernels  $\mathfrak{f}^u(\omega, \tau_1, \tau_2)$  defined in (5). To prove that result, we will use the following two regularity assumptions on  $X_{t, T}$ .

**(A)** There exists  $K < \infty$  such that for all  $t_0 \in \mathbb{Z}, T \in \mathbb{N}$  and  $(\tau_1, \tau_2) \in (0, 1)^2$ ,

$$\sum_{s \in \mathbb{Z}} |\gamma_{t_0; T}(s, \tau_1, \tau_2)| \leq K.$$

**(C)** For each  $t \in \mathbb{Z}$  and  $T \in \mathbb{N}$ , the function  $x \mapsto F_{t; T}(x)$  is continuous.

Assumption **(A)** is a uniform short-range dependence assumption. It holds under typical assumptions on short-range dependency, e.g. if the triangular array  $\{X_{t, T}\}$  is  $\alpha$ -mixing with mixing coefficients  $\alpha(k) \leq Ck^{-\delta}$  for some  $\delta > 1$  and  $C$  independent of  $T$ . Assumption **(C)** ensures that the marginal distributions  $F_{t; T}$  are continuous in a suitable uniform sense.

**Proposition 1.** Let  $\{X_{t,T}\}$  be locally strictly stationary, with approximating processes  $\{X_t(u)\}$ . If moreover  $\{X_{t,T}\}$  satisfies **(A)** and **(C)**, then

(a) the  $\gamma_h^u(\tau_1, \tau_2)$ 's are absolutely summable for any  $u$  and  $(\tau_1, \tau_2) \in (0, 1)^2$ , and

(b) for  $t_0 = \lfloor uT \rfloor$  and any  $(\tau_1, \tau_2) \in (0, 1)^2$ ,

$$\sup_{\omega \in (-\pi, \pi]} \left| f^u(\omega, \tau_1, \tau_2) - \mathfrak{W}_{t_0, T}(\omega, \tau_1, \tau_2) \right| = o(1)$$

as  $T \rightarrow \infty$ .

**Proof.** Throughout this proof, write  $t_0 = \lfloor uT \rfloor$ . Let  $G^u$  and  $F_{t,T}$  denote the marginal distribution functions of  $X_u(t)$  and  $X_{t,T}$  respectively. We begin by proving that under condition **(C)** the function  $x \mapsto G^u(x)$  is continuous for every  $u \in (0, 1)$ . This follows since, for any  $x \in \mathbb{R}$  and  $T \in \mathbb{N}$ ,

$$\lim_{y \rightarrow x} |G^u(x) - G^u(y)| \leq 4L/T + \lim_{y \rightarrow x} |F_{t_0;T}(x) - F_{t_0;T}(y)| = 4L/T.$$

This can be made arbitrarily small by choosing  $T$  sufficiently large, and continuity of  $G^u$  follows.

Next, observe that expectation of indicators can be written in terms of distribution functions, so that

$$\mathbb{E}(\mathbb{I}_{\{X_u(t) \leq x\}}) = G^u(x) \quad \text{and} \quad \mathbb{E}(\mathbb{I}_{\{X_u(t) \leq x\}} \mathbb{I}_{\{X_u(t-h) \leq y\}}) = G_h^u(x, y),$$

and therefore

$$\gamma_h^u(\tau_1, \tau_2) = G_h^u(q^u(\tau_1), q^u(\tau_2)) - \tau_1 \tau_2.$$

Using the same representation for  $\gamma_{t_0;T}(h, \tau_1, \tau_2)$ , we obtain

$$|\gamma_h^u(\tau_1, \tau_2) - \gamma_{t_0;T}(h, \tau_1, \tau_2)| \\ = \left| F_{[t_0-h/2], [t_0+h/2]; T}(F_{[t_0-h/2]; T}^{-1}(\tau_1), F_{[t_0+h/2]; T}^{-1}(\tau_2)) - G_h^u(q^u(\tau_1), q^u(\tau_2)) \right|.$$

Adding and subtracting  $G_h^u(F_{[t_0-h/2]; T}^{-1}(\tau_1), F_{[t_0+h/2]; T}^{-1}(\tau_2))$ , and utilizing the triangle inequality yields

$$\|F_{[t_0-h/2], [t_0+h/2]; T}(\cdot, \cdot) - G_h^u(\cdot, \cdot)\|_\infty \\ + \left| G_h^u(F_{[t_0-h/2]; T}^{-1}(\tau_1), F_{[t_0+h/2]; T}^{-1}(\tau_2)) - G_h^u(q^u(\tau_1), q^u(\tau_2)) \right|$$

where the first term, in view of local strict stationarity, can be bounded by

$$\|F_{[t_0-h/2], [t_0+h/2]; T}(\cdot, \cdot) - G_h^u(\cdot, \cdot)\|_\infty \leq L \left( \max \left( \left| \frac{t_0+h}{T} - u \right|, \left| \frac{t_0}{T} - u \right| \right) + 1/T \right) \\ \leq \frac{L(|h| + 2)}{T}.$$

For the second term, invoking Sklar's Theorem and the continuity of  $G^u$ , we can write

$$G_h^u(x, y) = \mathcal{C}(G^u(x), (G^u(y)))$$

where  $\mathcal{C}$  denotes the copula of  $G_h^u$ . As copulas are Lipschitz-continuous with constant one, we obtain

$$\begin{aligned} & |G_h^u(F_{[t_0-h/2];T}^{-1}(\tau_1), F_{[t_0+h/2];T}^{-1}(\tau_2)) - G_h^u(q^u(\tau_1), q^u(\tau_2))| \\ & \leq |G^u(F_{[t_0-h/2];T}^{-1}(\tau_1)) - G^u(q^u(\tau_1))| + |G^u(F_{[t_0+h/2];T}^{-1}(\tau_2)) - G^u(q^u(\tau_2))|. \end{aligned} \quad (7)$$

The continuity of  $G^u$  and  $F_{[t_0-h/2];T}$  implies, for the first term in the right-hand side of (7),

$$\begin{aligned} & |G^u(F_{[t_0-h/2];T}^{-1}(\tau_1)) - G^u(q^u(\tau_1))| = |G^u(F_{[t_0-h/2];T}^{-1}(\tau_1)) - \tau_1| \\ & = |G^u(F_{[t_0-h/2];T}^{-1}(\tau_1)) - F_{[t_0-h/2];T}(F_{[t_0-h/2];T}^{-1}(\tau_1))| \\ & \leq \|F_{[t_0-h/2];T}(\cdot) - G^u(\cdot)\|_\infty \leq L(|h| + 2)/T, \end{aligned}$$

where the last inequality holds due to local strict stationarity. Using the same arguments for the second term, we get

$$|\gamma_h^u(\tau_1, \tau_2) - \gamma_{t_0;T}(h, \tau_1, \tau_2)| \leq \frac{3L(|h| + 2)}{T}. \quad (8)$$

We first prove part (a) of the proposition. This will be done by contradiction. Assume that, for some  $u$ ,  $\tau_1$  and  $\tau_2$ , the  $\gamma_h^u(\tau_1, \tau_2)$ 's are not absolutely summable. In this case, there exists a  $H < \infty$  such that

$$\sum_{|h| \leq H} |\gamma_h^u(\tau_1, \tau_2)| > K + 1 \quad (9)$$

where  $K$  is the constant from assumption (A). On the other hand, given (8),

$$\sum_{|h| \leq H} |\gamma_h^u(\tau_1, \tau_2)| \leq \frac{3L}{T} \sum_{|h| \leq H} (|h| + 2) + \sum_{s \in \mathbb{Z}} |\gamma_{t_0;T}(s, \tau_1, \tau_2)| \leq \frac{3L}{T} (2 + 5H + H^2) + K.$$

By choosing  $T$  sufficiently large, this leads to a contradiction with (9). Absolute summability of  $\gamma_h^u(\tau_1, \tau_2)$  follows, hence part (a) of Proposition 1.

We now proceed to prove part (b). From the absolute summability of the  $\gamma_h^u(\tau_1, \tau_2)$ 's, we obtain

$$\mathfrak{f}^u(\omega, \tau_1, \tau_2) = \frac{1}{2\pi} \sum_{h=-T^{1/3}}^{T^{1/3}} \gamma_h^u(\tau_1, \tau_2) e^{-i\omega h} + o(1)$$

uniformly in  $\omega$ , while Assumption (A) yields, still uniformly in  $\omega$ ,

$$\mathfrak{W}_{t_0,T}(\omega, \tau_1, \tau_2) = \frac{1}{2\pi} \sum_{h=-T^{1/3}}^{T^{1/3}} \gamma_{t_0;T}(h, \tau_1, \tau_2) e^{-i\omega h} + o(1).$$

As a consequence of (8), we have

$$\sum_{h=-T^{1/3}}^{T^{1/3}} |\gamma_h^u(\tau_1, \tau_2) - \gamma_{t_0;T}(h, \tau_1, \tau_2)| = O\left(\frac{T^{2/3}}{T}\right) = o(1),$$

which establishes the desired result.  $\square$

The unicity of  $f^u(\omega, \tau_1, \tau_2)$  then follows as an immediate corollary.

**Corollary 1.** *For any locally strictly stationary  $\{X_{t,T}\}$  fulfilling (A) and (C), the time-varying copula spectral density  $f^u(\omega, \tau_1, \tau_2)$  is uniquely defined, i.e. does not depend of the choice of the approximating processes  $\{X_t(u)\}$ .*

## References

- [1] Birr, S., Volgushev, S., Kley, T., Dette, H., and Hallin, M. (2014). Quantile spectral analysis for locally stationary time series. *arXiv preprint arXiv:1404.4605*.
- [2] Claasen, T. A. C. M. and Mecklenbräuker, W. F. G. (1980). The Wigner distribution: A tool for time frequency signal analysis. *Philips J. Res.*, 35:217–250, 276–300, 372–389.
- [3] Dahlhaus, R. (1996). On the Kullback-Leibler information divergence of locally stationary processes. *Stochastic Processes and their Applications*, 62:139–168.
- [4] Dahlhaus, R. (1997). Fitting time series models to nonstationary processes. *Annals of Statistics*, 25:1–37.
- [5] Davis, R. A., Mikosch, T. and Zhao, Y. (2013). Measures of serial extremal dependence and their estimation. *Stochastic Processes and their Applications*, 123:2575–2602.
- [6] Dette, H., Hallin, M., Kley, T., and Volgushev, S. (2015). Of copulas, quantiles, ranks and spectra: An  $L_1$  approach to spectral analysis. *Bernoulli*, 21:781–831.
- [7] Flandrin, P. and Escudid, B. (1980). Time and frequency representation of finite energy signals a physical property as a result of an Hilbertian condition. *Signal Processing*, 2:93–100.
- [8] Hagemann, A. (2013). Robust spectral analysis. Available at arXiv:1111.1965v2.
- [9] Hlawatsch, F. and Boudreux-Bartels, G. F. (1992). Linear and quadratic time-frequency signal representations. *IEEE Signal Processing Magazine*, 9:21–67.

- [10] Hong, Y. (1999). Hypothesis testing in time series via the empirical characteristic function: A generalized spectral density approach. *Journal of the American Statistical Association*, 94:1201–1220.
- [11] Kley, T., Volgushev, S., Dette, H., and Hallin, M. (2016). Quantile spectral processes: Asymptotic analysis and inference. *Bernoulli*, 22:1770–1807.
- [12] Kley, T. (2016). Quantile-Based Spectral Analysis in an Object-Oriented Framework and a Reference Implementation in R: The quantspec Package. *Journal of Statistical Software*, 70(3):1–27.
- [13] Lee, J. and Subba Rao, S. (2012). The quantile spectral density and comparison-based tests for nonlinear time series. Available at arXiv:1112.2759v2.
- [14] Li, T.H. (2008). Laplace periodogram for time series analysis. *Journal of the American Statistical Association*, 103:757–768.
- [15] Li, T.H. (2012). Quantile periodograms. *Journal of the American Statistical Association*, 107:765–776.
- [16] Li, T.H. (2014). Quantile periodogram and time-dependent variance. *Journal of Time Series Analysis*, 35:322–340.
- [17] Loynes, R. M. (1968). On the Concept of the Spectrum for Non-Stationary Processes. *Journal of the Royal Statistical Society Ser. B*, 1:1–30.
- [18] Mark, W. D. (1970). Spectral analysis of the convolution and filtering of non-stationary stochastic processes. *Journal of Sound and Vibration*, 11(1):19–63.
- [19] Martin, W. (1982). Time-frequency analysis of random signals. In *Acoustics, Speech, and Signal Processing, IEEE International Conference on ICASSP'82.*, volume 7, pages 1325–1328. IEEE.
- [20] Martin, W. and Flandrin, P. (1985). Wigner-Ville spectral analysis of nonstationary processes. *IEEE Transactions on Acoustics, Speech, and Signal Processing*, 33:1461–1470.
- [21] Parzen, E. (2004). Quantile probability and statistical data modeling. *Statistical Science*, 19:652–662.
- [22] Priestley, M. B. (1965). Evolutionary spectra and non-stationary processes. *Journal of the Royal Statistical Society Ser. B*, 27:204–237.
- [23] Roueff, F. and Von Sachs, R. (2011). Locally stationary long memory estimation. *Stochastic Processes and their Applications*, 121: 813–844.
- [24] Ville, J. (1948). Théorie et applications de la notion de signal analytique. *Cables et Transmissions*, 2A:61–74.

- [25] Vogt, M. (2012). Nonparametric regression for locally stationary time series. *Annals of Statistics*, 40:2601–2633.
- [26] Wigner, E. (1932). On the quantum correction for thermodynamic equilibrium. *Physical Review*, 40:749–759.
- [27] Zhou, Z. and Wu, W.B. (2009a). Local linear quantile estimation for non-stationary time series. *Annals of Statistics*, 37:2696–2729.
- [28] Zhou, Z. and Wu, W.B. (2009b). Nonparametric inference of discretely sampled stable Lévy processes. *Annals of Statistics*, 37:83–92.